

5010 solutions, Assignment 9. Chapter 5: 31, 32, 33, 35, 36, 39.

31. Let  $X_1, X_2, \dots, X_k$  be the numbers on the removed balls.

(a)  $E[X_1 + \dots + X_k] = kE[X_1] = k(1 + 2 + \dots + n)/n = k(n+1)/2$ , and  $\text{Var}(X_1 + \dots + X_k) = k\text{Var}(X_1) + k(k-1)\text{Cov}(X_1, X_2)$ . Now  $\text{Var}(X_1) = (1^2 + 2^2 + \dots + n^2)/n - [(1+2+\dots+n)/n]^2 = (n+1)(2n+1)/6 - (n+1)^2/4 = (n^2-1)/12$  and  $\text{Cov}(X_1, X_2) = \sum_{i \neq j} ij/[n(n-1)] - (n+1)^2/4 = (\sum_{i,j} ij - \sum_i i^2)/[n(n-1)] - (n+1)^2/4$ , etc., etc.

(b) Mean is the same as in (a), variance is  $k(n^2-1)/12$  from (a) since the covariance terms are 0.

(c) Let  $M$  be the largest number removed. In (a),

$$P(M = m) = \frac{\binom{m}{k} \binom{n-m}{0}}{\binom{n}{k}} - \frac{\binom{m-1}{k} \binom{n-m+1}{0}}{\binom{n}{k}} = \frac{\binom{m-1}{k-1}}{\binom{n}{k}}.$$

and in (b),

$$P(M = m) = (m/n)^k - ((m-1)/n)^k,$$

since in both case you have to choose all balls from  $\{1, 2, \dots, m\}$  but not from  $\{1, 2, \dots, m-1\}$ .

32.  $X - Y$  and  $X + Y$  assume the values  $\pm a$  with probability  $1/2$  each, and we can check that, and the joint distribution assumes all four possibilities  $(a, a), (a, -a), (-a, a), (-a, -a)$  with probability  $1/4$  each. Thus, the product of the marginal distributions is the joint distribution, which is independence.

33. The joint distribution is  $P(U = 1, V = 1) = P(U = 1)P(V = 1 | U = 1) = (1/2)(1/3) = 1/6$ ,  $P(U = 1, V = -1) = P(U = 1)P(V = -1 | U = 1) = (1/2)(2/3) = 1/3$ ,  $P(U = -1, V = 1) = P(U = -1)P(V = 1 | U = -1) = (1/2)(2/3) = 1/3$ , and  $P(U = -1, V = -1) = P(U = -1)P(V = -1 | U = -1) = (1/2)(1/3) = 1/6$ .

(a) The condition for real roots is that the discriminant is nonnegative, i.e.,  $U^2 - 4V \geq 0$ . Since  $U$  and  $V$  are  $\pm 1$ , it is necessary and sufficient that  $V = -1$ , and so the probability is  $1/2$ .

(b)  $P(U = 1 | V = -1) = (1/3)/(1/2) = 2/3$ , hence  $P(U = -1 | V = -1) = 1/3$ . If  $V = -1$ , the largest root of the quadratic is  $(-U + \sqrt{U^2 + 4})/2 = (-U + \sqrt{5})/2$ . Its expectation is  $(2/3)(-1 + \sqrt{5})/2 + (1/3)(1 + \sqrt{5})/2 = -1/6 + \sqrt{5}/2$ .

(c)  $U + V$  equals 2 with probability  $1/6$ , 0 with probability  $2/3$ , and  $-2$  with probability  $1/6$ . The discriminant is  $(U + V)^2 - 4(U + V) = (U + V - 4)(U + V)$ . This will be nonnegative if  $U + V$  is 0 or  $-2$ , hence with probability  $5/6$ .

35. (a) Let  $Y$  be the number requiring surgery. Then

$$\begin{aligned} P(Y = m) &= \sum_{n \geq m} P(X = n)P(Y = m | X = n) \\ &= \sum_{n \geq m} e^{-8} 8^n (n!)^{-1} \binom{n}{m} (1/4)^m (3/4)^{n-m} \end{aligned}$$

$$\begin{aligned}
&= e^{-8} 3^{-m} (m!)^{-1} \sum_{n \geq m} 8^n (3/4)^n / (n-m)! \\
&= e^{-8} 3^{-m} (m!)^{-1} 6^m e^6 \\
&= e^{-2} 2^{-m} (m!)^{-1},
\end{aligned}$$

which is Poisson(2), whose mean is 2.

(b) Replace 8 by 4 to get Poisson(1) for weekends. For the week, we have 5 days with Poisson(2) and 2 days of Poisson(1). Total is Poisson(12), which has mean and variance 12.

36. This is the hypergeometric distribution. The mean is  $E[X_1 + \dots + X_n] = nE[X_1] = nm/M$ , and the variance is  $n\text{Var}(X_1) + n(n-1)\text{Cov}(X_1, X_2) = n(m/M)(1-m/M) + n(n-1)\{m(m-1)/[M(M-1)] - (m/M)^2\}$ , which can be simplified slightly.

39. (a)

$$\begin{aligned}
P(X_1 < X_2 < X_3) &= \sum_{i=1}^{\infty} \sum_{j=i+1}^{\infty} \sum_{k=j+1}^{\infty} P(X_1 = i, X_2 = j, X_3 = k) \\
&= \sum_{i=1}^{\infty} \sum_{j=i+1}^{\infty} \sum_{k=j+1}^{\infty} p_1^{i-1} (1-p_1) p_2^{j-1} (1-p_2) p_3^{k-1} (1-p_3) \\
&= \sum_{i=1}^{\infty} \sum_{j=i+1}^{\infty} p_1^{i-1} (1-p_1) p_2^{j-1} (1-p_2) p_3^j \\
&= \sum_{i=1}^{\infty} \sum_{j=i+1}^{\infty} p_1^{i-1} (1-p_1) (p_2 p_3)^{j-1} (1-p_2) p_3 \\
&= \sum_{i=1}^{\infty} p_1^{i-1} (1-p_1) (p_2 p_3)^i (1-p_2) p_3 / (1-p_2 p_3) \\
&= \sum_{i=1}^{\infty} (p_1 p_2 p_3)^{i-1} (1-p_1) (1-p_2) p_2 p_3^2 / (1-p_2 p_3) \\
&= (1-p_1) (1-p_2) p_2 p_3^2 / [(1-p_2 p_3) (1-p_1 p_2 p_3)],
\end{aligned}$$

(b)

$$\begin{aligned}
P(X_1 \leq X_2 \leq X_3) &= \sum_{i=1}^{\infty} \sum_{j=i}^{\infty} \sum_{k=j}^{\infty} P(X_1 = i, X_2 = j, X_3 = k) \\
&= \sum_{i=1}^{\infty} \sum_{j=i}^{\infty} \sum_{k=j}^{\infty} p_1^{i-1} (1-p_1) p_2^{j-1} (1-p_2) p_3^{k-1} (1-p_3) \\
&= \sum_{i=1}^{\infty} \sum_{j=i}^{\infty} p_1^{i-1} (1-p_1) p_2^{j-1} (1-p_2) p_3^{j-1}
\end{aligned}$$

$$\begin{aligned}
&= \sum_{i=1}^{\infty} \sum_{j=i}^{\infty} p_1^{i-1} (1-p_1) (p_2 p_3)^{j-1} (1-p_2) \\
&= \sum_{i=1}^{\infty} p_1^{i-1} (1-p_1) (p_2 p_3)^{i-1} (1-p_2) / (1-p_2 p_3) \\
&= \sum_{i=1}^{\infty} (p_1 p_2 p_3)^{i-1} (1-p_1) (1-p_2) / (1-p_2 p_3) \\
&= (1-p_1) (1-p_2) / [(1-p_2 p_3) (1-p_1 p_2 p_3)],
\end{aligned}$$

(c) The probability A throws 6 first, B second, and C third is just the answer to part (b) with  $p_1 = p_2 = p_3 = 5/6$ . We get

$$(1/6)(1/6) / [(1 - 25/36)(1 - 125/216)] = 216/1001.$$